Fitting generalized estimating equation (GEE) regression models in Stata

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Outline

- Regression models for clustered or longitudinal data
- Brief review of GEEs
  - mean model
  - working correlation matrix
- Stata GEE implementation
- Example: Mental health service utilization
- Summary and conclusions

Regression models for clustered or longitudinal data

- Longitudinal, repeated measures, or clustered data commonly encountered
- Correlations between observations on a given subject may exist, and need to be accounted for
- If outcomes are multivariate normal, then established methods of analysis are available (Laird and Ware, Biometrics, 1982)
- If outcomes are binary or counts, likelihood based inference less tractable

Generalized estimating equations

- Described by Liang and Zeger (Biometrika, 1986) and Zeger and Liang (Biometrics, 1986) to extend the generalized linear model to allow for correlated observations
- Characterize the marginal expectation (average response for observations sharing the same covariates) as a function of covariates
- Method accounts for the correlation between observations in generalized linear regression models by use of empirical (sandwich/robust) variance estimator
- Posits model for the working correlation matrix

The marginal mean model

- We assume the marginal regression model:

$$g(E[Y_{ij} \mid x_{ij}]) = x_{ij}^{'}\beta$$

- Where x_{ij} is a p times 1 vector of covariates, β consists of the p regression parameters of interest, g(.) is the link function, and Y_{ij} denotes the *j*th outcome (for j=1,…,J) for the *i*th subject (for i=1,…,N)
- Common choices for the link function include:

  g(a)=a (identity link)

  g(a)=log(a) [for count data]

  g(a)=log(a/(1-a)) [logit link for binary data]

Model for the correlation

- Assuming no missing data, the J x J covariance matrix for Y is modeled as:

$$V_i = \phi A_i^{1/2} R(\alpha) A_i^{1/2}$$

- Where ϕ is a glm dispersion parameter, A is a diagonal matrix of variance functions, and $R(\alpha)$ is the working correlation matrix of Y

Model for the correlation (cont.)

- If mean model is correct, correlation structure may be mis-specified, but parameter estimates remain consistent
- Liang and Zeger showed that modeling correlation may boost efficiency
- But this is a large sample result; there must be enough clusters to estimate these parameters
- Variety of models that are supported in Stata

Model for the correlation (cont.)

- Independence

$$R(\alpha) = \begin{pmatrix} 1 & 0 & \cdots & 0 \\ 0 & 1 & \cdots & 0 \\ \vdots & \vdots & \ddots & \vdots \\ 0 & 0 & \cdots & 1 \end{pmatrix}$$

- Number of parameters: 0

Model for the correlation (cont.)

- Exchangeable (compound symmetry)

$$R(\alpha) = \begin{pmatrix} 1 & \alpha & \cdots & \alpha \\ \alpha & 1 & \cdots & \alpha \\ \vdots & \vdots & \ddots & \vdots \\ \alpha & \alpha & \cdots & 1 \end{pmatrix}$$

- Number of parameters: 1

Model for the correlation (cont.)

- Unstructured

$$R(\alpha) = \begin{pmatrix} 1 & \alpha_{12} & \cdots & \alpha_{1J} \\ \alpha_{12} & 1 & \cdots & \alpha_{2J} \\ \vdots & \vdots & \ddots & \vdots \\ \alpha_{1J} & \alpha_{2J} & \cdots & 1 \end{pmatrix}$$

- Number of parameters: J(J-1)/2

Model for the correlation (cont.)

- Auto-regressive

$$R(\alpha) = \begin{pmatrix} 1 & \alpha & \cdots & \alpha^{J-1} \\ \alpha & 1 & \cdots & \alpha^{J-2} \\ \vdots & \vdots & \ddots & \vdots \\ \alpha^{J-1} & \alpha^{J-2} & \cdots & 1 \end{pmatrix}$$

- Number of parameters: 1

Model for the correlation (cont.)

- Stationary (g-dependent)

$$R(\alpha) = \begin{pmatrix} 1 & \alpha_1 & \cdots & \alpha_{J-1} \\ \alpha_1 & 1 & \cdots & \alpha_{J-2} \\ \vdots & \vdots & \ddots & \vdots \\ \alpha_{J-1} & \alpha_{J-2} & \cdots & 1 \end{pmatrix}$$

- Number of parameters: 0 <g <= J-1

Model for the correlation (cont.)

- Fixed

$$R(\alpha) = \begin{pmatrix} 1 & c_{12} & \cdots & c_{1J} \\ c_{12} & 1 & \cdots & c_{2J} \\ \vdots & \vdots & \ddots & \vdots \\ c_{1J} & c_{2J} & \cdots & 1 \end{pmatrix}$$

- Number of parameters: 0 (user specified)

Model for the correlation (cont.)

- If J is small and data are balanced and complete, then an unstructured matrix is recommended
- If observations are mistimed, then use a structure that accounts for correlation as function of time (stationary, or auto-regressive)
- If observations are clustered (i.e. no logical ordering) then exchangeable may be appropriate
- If number of clusters small, independent may be best
- Issues discussed further in Diggle, Liang and Zeger (1994, book)

Missing data

- Standard GEE models assume that missing observations are Missing Completely at Random (MCAR) in the sense of Little and Rubin (book, 1987)
- Robins, Rotnitzky and Zhao (JASA, 1995) proposed methods to allow for data that is missing at random (MAR)
- These methods not yet implemented in standard software (requires estimation of weights and more complicated variance formula)

Variance estimators

- Empirical (aka *sandwich* or *robust/semi-robust*)

  consistent when the mean model is correctly specified (if no missing data)
- Model-based (aka *naïve*) [default in Stata]

  consistent when both the mean model and the covariance model are correctly specified

Syntax for xtgee

xtgee depvar varlist, family(family) link(link) corr(corr) i(idvar) t(timevar) robust

Family: binomial, gaussian, gamma, igaussian, nbinomial, poisson

Link: identity, cloglog, log, logit, nbinomial, opwer, power, probit, reciprocal

Correlation: independent, exchangeable, ar#, stationary#, nonstationary#,unstructured, fixed

Also options to change the scale parameter, use weighted equations, specify offsets

Example: Mental Health Service Utilization

- Connecticut child studies (Zahner et al, AJPH, 1997)
- Outcome: use of general health, school, or mental health services (dichotomous report)
- Sample: 2,519 children
- Other dichotomous predictors: age, gender, academic problems

Data format and variables

OBS	ID	BOY	OLD	ACADPR	SETTING	SCHOOL	MENTAL	GENERAL	SERV
1	90111502	0	0	0	0	1	0	0	0
2	90111502	0	0	0	1	0	1	0	0
3	90111502	0	0	0	2	0	0	1	0
4	**80111206**	**0**	**0**	**0**	**0**	**1**	**0**	**0**	**0**
5	**80111206**	**0**	**0**	**0**	**1**	**0**	**1**	**0**	**0**
6	**80111206**	**0**	**0**	**0**	**2**	**0**	**0**	**1**	**0**
7	40111608	1	0	0	0	1	0	0	0
8	40111608	1	0	0	1	0	1	0	0
9	40111608	1	0	0	2	0	0	1	0

Stata code to fit model

```
iis id
tis setting
xtdes
xi: xtgee serv i.old*mental i.old*school
  i.boy*mental i.boy*school
  i.acadpro*mental i.acadpro*school,
  link(logit) corr(unst) family(binomial)
  robust
xtcorr
```

Describe cross-sectional data (xtdes)

```
id: 1, 2, ..., 2519                         n =       2519
setting:  0, 1, ..., 2                      T =          3
          Delta(type) = 1; (2-0)+1 = 3
          (id*setting uniquely identifies each observation)

Distribution of T_i: min    5%   25%     50%     75%   95%   max
                       3     3     3       3       3     3     3

     Freq.  Percent    Cum. |  Pattern
 ---------------------------+---------
     2519    100.00  100.00 |  111
 ---------------------------+---------
     2519    100.00         |  XXX
```

(No missing data!)


```
GEE population-averaged model                 Number of obs      =      7557
Group and time vars:          id setting      Number of groups   =      2519
Link:                              logit      Obs per group: min =         3
Family:                         binomial                     avg =       3.0
Correlation:                unstructured                     max =         3
                                            Wald chi2(11)      =    605.12
Scale parameter:                         1  Prob > chi2        =    0.0000
                       (standard errors adjusted for clustering on id)
------------------------------------------------------------------------------
             |             Semi-robust
        serv |      Coef.   Std. Err.      z    P>|z|   [95% Conf. Interval]
-------------+----------------------------------------------------------------
     _Iold_1 |   .1233576   .1441123     0.86   0.392    -.1590973  .4058124
      mental |  -.3520988   .1933698    -1.82   0.069    -.7310967  .0268992
_IoldXment~1 |   .2905076    .189558     1.53   0.125    -.0810192  .6620344
      school |   .1850487   .1734874     1.07   0.286    -.1549804  .5250778
_IoldXscho~1 |    .330549    .162133     2.04   0.041     .0127742  .6483239
     _Iboy_1 |   .3652564   .1464068     2.49   0.013     .0783043  .6522084
_IboyXment~1 |  -.2779134   .1894824    -1.47   0.142    -.6492921  .0934654
_IboyXscho~1 |  -.1538587   .1650033    -0.93   0.351    -.4772592  .1695418
 _Iacadpro_1 |   .7239641   .1445971     5.01   0.000      .440559  1.007369
_IacaXment~1 |   .1843236   .1911094     0.96   0.335    -.1902441  .5588912
_IacaXscho~1 |   1.136088   .1669423     6.81   0.000     .8088873  1.463289
       _cons |  -2.944382   .1489399   -19.77   0.000    -3.236298 -2.652465
```

Estimates of working correlation (xtcorr)

```
Estimated within-id corr matrix R

     school mental general
         c1     c2     c3
r1   1.0000
r2   0.1646  1.0000
r3   0.1977  0.2270  1.0000
```


Multidimensional test of OLD effect

```
test _IoldXmenta_1=0
( 1)  _IoldXmenta_1 = 0.0
      chi2(  1) =    2.35
      Prob > chi2 =  0.1254
test _IoldXschoo_1=0,accumulate
( 1)  _IoldXschoo_1 = 0.0
( 2)  _IoldXmenta_1 = 0.0
      chi2(  2) =    4.55
      Prob > chi2 =  0.1029     ←
test _Iold_1=0,accumulate
( 1)  _IoldXschoo_1 = 0.0
( 2)  _IoldXmenta_1 = 0.0
( 3)  _Iold_1 = 0.0
      chi2(  3) =   20.61
      Prob > chi2 =  0.0001     ←
```

Results from Example

- There is a significant interaction between service setting and academic problems (df=2,p<0.0001), but not for age and setting (df=2,p=0.10) or gender and setting (df=2,p=0.33)
- Overall, a higher proportion of boys use services (df=3,p=0.04) and older children use them more than younger children (df=3,p=0.0001)

More resources

- Generalized estimating equations: an annotated bibliography (Ziegler, Kastner and Blettner, Biometrical Journal, 1998)
- Review of software to fit Generalized Estimating Equation regression models (Horton and Lipsitz, The American Statistician, 1999, article online at http://www.biostat.harvard.edu/~horton/geereview.pdf)